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Aid, Dutch disease, and manufacturing growth

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ABSTRACT

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1. Introduction

There is a voluminous literature on the effect of foreign aid on economic growth.¹ A fair summary would be that this literature does not provide robust evidence of either a positive or negative correlation between foreign aid inflows and the economic growth of poor countries. The absence of a robust positive correlation is particularly surprising since the primary purpose of most aid flows is to further the economic development of the recipient. What might explain the paucity of evidence for foreign aid propelling economic growth?

One possible answer may be that there are costs emanating from aid that offsets the benefits—which include higher levels of physical investment, better education, and health, all paid for by the resource transfer. The classic analysis of these costs goes back to the Keynes— Ohlin debates on the effects of foreign transfers (such as aid) on the real exchange rate. For the case of foreign aid, the seminal work is van

lower relative growth rate of exportable industries. We provide some evidence suggesting that the channel for these effects is the real exchange rate appreciation caused by aid inflows. We conjecture that this may explain, in part, why it is hard to find robust evidence that foreign aid helps countries grow. © 2009 Elsevier B.V. All rights reserved.

We examine the effects of aid on the growth of manufacturing, using a methodology that exploits the

variation within countries and across manufacturing sectors, and corrects for possible reverse causality.

We find that aid inflows have systematic adverse effects on a country's competitiveness, as reflected in the

Wijnbergen (1985, 1986) building on the work of Corden and Neary (1982) in the context of natural resource discoveries.

In a simple two-good model, increased aid will have two effects. First, aid could disproportionately be targeted at expanding non-tradable services such as construction, health care, and education for which there is substantial unmet demand. This will increase wages in that sector (given a fixed supply of skilled labor in the short run), will draw skilled labor into the non-tradable sector, and will increase wages overall. Given that the international price of traded goods is fixed, the higher wage in terms of traded goods will reduce traded sector profitability, competitiveness, and lead to a decline in exports. In Corden and Neary (1982)'s terminology, this is the *resource movement* effect.

In addition, the higher wages will be spent, raising the price of non-traded goods relative to traded goods (the real exchange rate) and further hurting traded sector competitiveness. Corden and Neary (1982) would term this the *spending* effect.

These are short run effects. But, in principle, in the medium run, aid could simultaneously also increase the productivity and supply of non-traded goods (for example, if the aid is spent on imparting skills and expanding the availability of skilled labor), offsetting the initial spending and resource movement impacts on relative prices. The overall impact, at least in the medium run—which is the focus of this paper—cannot be predicted a priori (also see Torvik (2001)). The more aid is spent on traded goods or factors (imported capital goods and foreign consultants) and on factors that are not in limited supply (unskilled labor), the more the supply of factors and non-traded

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¹ Important work includes Boone (1996), Alesina and Weder (2002), Bauer (1971), Burnside and Dollar (2000), Clemens, Radelet, and Bhavnani (2004), Collier and Dollar (2002), Dalgaard, Hansen, and Tarp (2004), Easterly (2003), Easterly, Levine, and Roodman (2004), Friedman (1958), Hansen and Tarp (2001), Rajan and Subramanian (2008), Roodman (2004), Svensson (2003), and World Bank (1998).

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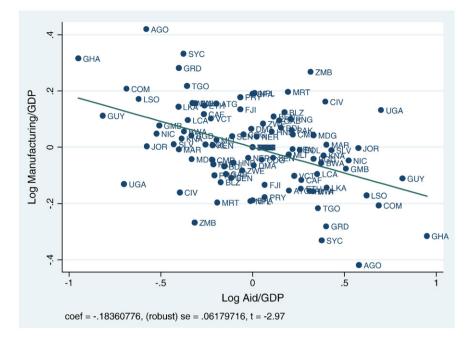


Fig. 1. Manufacturing and aid between 1980 and 2000. This plot is based on running a panel regression where the dependent variable is the log of the ratio of the share of value added in manufacturing to GDP for a country (at two dates, the late 1990s and the early 1980s), and the explanatory variables are the country's per capita PPP GDP, per capita PPP GDP squared, and fixed effects for the country and the time period. It can be interpreted as representing the conditional relationship between the change in the size of the manufacturing sector between 1980 and 2000 in a country and the change in aid over the same period. All variables are averages for the period 1980–85 and 1995–2000, respectively. To facilitate comparability with the core results in the paper, the sample was chosen according to the same criteria as in the core sample of the paper, namely to include countries that had an aid-to-GDP ratio greater than 1% or are low-income countries. Data on manufacturing are from the World Bank's World Development Indicators.

goods respond to aid inflows, and the more domestic fiscal contraction takes place, the less likely will wages and prices be bid up to an excessive degree and the less likely will the real exchange rate appreciate (see Berg et al. (2005)). Thus whether aid reduces competitiveness and shrinks the tradable goods sector by pushing up the real exchange rate is ultimately an empirical question.

A hint that aid may have costs via its impact on the tradable sector is provided in Fig. 1, where we plot the log of the value added in manufacturing to GDP ratio in a country against the log of the ratio of aid received to GDP for that country for two separate dates (the late 1990s and the early 1980s, separated by about 15 years), after correcting for the country's per capita PPP GDP, per capita PPP GDP squared, and country and time fixed effects.² As the figure suggests, the more aid a country has received, the smaller its share of manufacturing. The coefficient estimate suggests that a 1 percentage point increase in the ratio of aid-to-GDP is associated with a reduced share of manufacturing in total GDP of about 0.2–0.3 percentage points.³

In Fig. 2 we obtain a similar negative relationship when we plot the log of the ratio of manufacturing share to the share of services against the log of aid-to-GDP after similar corrections. So the inflow of aid is correlated with slower manufacturing growth, and the slower growth of manufacturing relative to services.

What might explain this pattern? Assuming that services are less tradable than manufacturing (at least for the period covered by our analysis), the figures suggest that aid and the relative size of tradable sectors in an economy are negatively correlated. The focus of our paper is then on providing more persuasive evidence for the channel through which this correlation emerges; that the inflow of aid leads to a relative shrinkage of the tradable manufacturing sectors of the economy and that this shrinkage occurs through the appreciation of the recipient country's real exchange rate.

Why might this matter? For instance, a number of studies (Jones and Olken (2005) and Rodrik (2007)) have argued that the traded goods sector is the channel through which an economy absorbs best practices from abroad. The absence of these *learning-by-doing spillovers*, which may be critical to long-run productivity growth, could be one constraint on growth. Indeed, Jones and Olken (2005) and Johnson, Ostry, and Subramanian (forthcoming) point out that virtually all countries that have had a sustained period of growth in the post-war period have seen a large increase in their share of manufacturing and manufacturing exports.

While these studies do not establish beyond doubt that there is a line of causation from manufacturing exports to overall growth—and this paper does not address this question—the existence of such a line would explain why the positive effects of aid are so hard to discern—aid leads to a shrinkage of the tradable manufacturing sector, which has adverse effects on overall growth.

There are three notable differences between our approach and that of prior empirical work. First, the papers in the literature on the effects of aid on growth of manufacturing or of manufacturing exports employ the standard cross-country methodology.⁴ The key innovation in our paper is to identify the effect of aid on manufacturing growth in a within-

² The residuals are obtained in a panel regression where the dependent variable is log of the ratio of the share of value added in manufacturing to GDP for a country (at two dates, the late 1990s and the early 1980s, separated by about 15 years), and the explanatory variables are the country's per capita GDP, per capita GDP squared (to allow for the U-shaped relationship postulated by Kuznets), and fixed effects for the country and the time period. All variables are averages for the period 1980–85 and 1995–2000, respectively. We focus on the period after 1980 because the coverage of the World Bank data on sectoral output increases substantially and also because the core analysis in the paper below is on the 1980s and 1990s. The sample comprises countries that had an aid received to GDP ratio greater than 1% or are low income as the terms of trade.

³ Because we include fixed effects, the association between aid and manufacturing depicted in the chart is a temporal one, that is, within countries over time rather than one between countries.

⁴ The empirical literature on aid, exchange rates, and manufacturing includes Adam (2005), Arellano et al. (2005), Atingi-Ego and Sebudde (2000), Bulir and Lane (2002), Elbadawi (1999), Nyoni (1998), Prati and Tressel (2006), Vos (1998), Yano and Nugent (1999) and Younger (1992) but none of them exploits within-sector variation. A fuller description of this literature is in the online version of this paper.

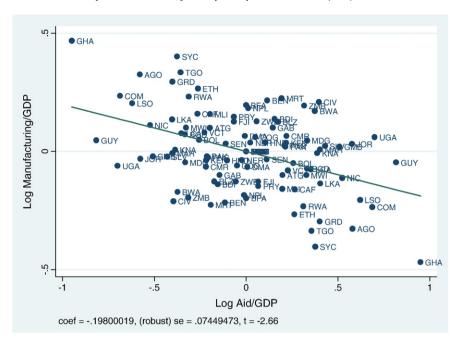


Fig. 2. Share of manufacturing relative to services and aid between 1980 and 2000. This plot represents the conditional relationship between the change in the size of the manufacturing sector relative to the size of the service sector between 1980 and 2000 in a country and the change in aid over the same period. The share of the service sector is obtained as 1-(share of agriculture + share of industry). It is based on running a panel regression where the dependent variable is the log of the ratio of the share of value added in manufacturing to GDP for a country (at two dates, the late 1990s and the early 1980s), and the explanatory variables are the country's per capita PPP GDP, per capita PPP GDP squared, and fixed effects for the country and the time period. All variables are averages for the paper, the sample was chosen according to the same criteria as in the core sample of the paper, namely, to include countries that had an aid-to-GDP ratio greater than 1% or are low-income countries. Data on manufacturing and services are from the World Bank's World Development Indicators.

country cross-industry context, using the methodology in Rajan and Zingales (1998). By absorbing country-specific variation in country indicators, and exploiting only the within-country across-industry variation, we go some way in addressing the specification problem that plague standard cross-country regressions—that some omitted country-specific variable might explain the observed correlation.

Second, the core of the paper aims to establish the effect of aid on the relative share of exportable sectors, and to provide evidence on the channel through which aid might have such effects. Specifically, we show that the effects of aid on relative industry growth rates seem to flow largely through the real exchange rates. In other words, we seek to establish that aid is the deep determinant, while exchange rates are the proximate transmission mechanism.

Third, countries may get more aid if they perform particularly badly so aid is clearly endogenous to growth (though less clearly linked to relative industry growth rates). We address the issue of endogeneity by using plausibly exogenous instruments for aid.

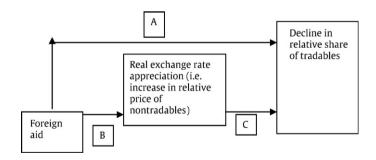
To preview our results, we find strong evidence that aid undermines the competitiveness of the exportable sectors. In particular, in countries that receive more aid, manufacturing industries that typically export more, and thus are more likely to be adversely affected by an appreciating real exchange rate, grow slower than manufacturing industries that typically export less.

If indeed the channel is through aid's effect on real exchange rates, we should see that: (1) aid inflows do affect exchange rates and lead to exchange rate appreciation in our sample; (2) in countries with a greater appreciation of the real exchange rate (or, more precisely, with a greater aid-induced appreciation) more exportable industries grow relatively more slowly; and (3) the independent effect of aid is somewhat attenuated when we include the extent to which the exchange rate has appreciated above the norm. We do see all these effects.

This paper is structured as follows. In Section 2, we outline our empirical strategy. In Section 3, we present our main results on the effect of aid on the relative growth of sectors whose competitiveness is most likely to be adversely hit, and establish its robustness. In Section 4, we provide evidence that exchange rate overvaluation (and not other policies or institutions) is the proximate channel through which aid has an effect, and Section 5 provides a conclusion.

2. Aid and growth: empirical strategy

The following is a schematic representation of the effects of aid on prices and quantities.



Simply put, we hypothesize a relationship from aid to the growth of the tradable sector (channel A). Channel A can in turn be broken down to two mediating channels: from aid to real exchange rates (B) and from real exchange rates to the growth of the tradable sector (C). We do not have a fully elaborated structural model that can test all these effects simultaneously so we will test for these individual relationships in the data.

We use the methodology in Rajan and Zingales (1998) to test the hypothesis that aid inflows might reduce the competitiveness of the traded goods sector. They suggest that one way to check whether a channel is at work is to see whether industries that might be most affected by a channel grow differentially (faster or slower depending

Summary statistics.

Variables	Period	Mean	Median	Standard deviation	Minimum	Maximum	Number of observations
A. Across countries and industries in the base sample							
Growth rate of value added ij	1980s	0.018	0.017	0.121	-0.457	0.754	684
	1990s	0.030	0.037	0.108	-0.530	0.337	357
Initial industry share ij	1980s	0.045	0.022	0.068	0.000	0.562	684
	1990s	0.041	0.020	0.064	0.000	0.525	357
B. Across industries in the base sample							
Exportability1 index i	1980s, 1990s	0.5	0	0.500	0	1	28
Exportability2 index i	1980s, 1990s	0.14	0	0.354	0	1	28
Financial dependence i	1980s	0.233	0.219	0.322	-0.451	1.140	27
C. Across countries in the base sample							
Aid to GDP j	1980s	0.053	0.036	0.039	0.008	0.176	32
	1990s	0.047	0.036	0.040	0.006	0.139	15
Average real exchange rate overvaluation j	1980s	-0.077	-0.153	0.355	-0.667	0.664	31
	1990s	-0.133	-0.011	0.368	-0.721	0.821	15
Policy (Sachs-Warner) index j	1980s	0.279	0.000	0.378	0.000	1.000	32
	1990s	0.775	1.000	0.335	0.000	1.000	15
Black market premium j	1980s	0.698	0.174	1.242	-0.008	5.936	30
	1990s	0.142	0.054	0.288	0.017	1.247	15
Initial per capita Income <i>j</i> (in log)	1980s	7.639	7.732	0.609	6.406	8.659	31
	1990s	7.750	7.830	0.708	6.202	9.106	15
Life expectancy j	1980s	56.53	54.83	7.04	44.21	72.70	32
	1990s	60.95	61.71	7.56	45.00	70.31	15
Geography j	1980s	-0.709	-1.015	0.552	-1.040	0.724	32
	1990s	-0.641	-1.014	0.604	-1.040	0.724	15
Institutional quality (ICRG) index j	1980s	0.409	0.399	0.127	0.183	0.723	28
	1990s	0.528	0.526	0.054	0.405	0.610	14
Average level of tariffs j	1980s	33.279	29.000	22.047	4.000	97.000	32
	1990s	23.277	24.000	8.291	10.000	45.000	15
Standard deviation of inflation j	1980s	153.021	6.348	712.080	2.053	3662.771	32
	1990s	6.553	4.691	4.043	1.631	15.855	15

on the nature of the effect) in countries where that channel is likely to be more operative. The industry characteristic we are interested in is the degree to which an industry's competitive position is affected by exchange rate appreciation, the channel is real exchange rate appreciation, and countries that get more aid are likely to be the ones where the channel is most operative. The estimation strategy is then to run regressions of the form:

 $Growth_{ii} = Constant + \zeta_{1,...,m}^*Country Indicators$

+ $\zeta_{m + 1...n}$ *Industry Indicators + $\zeta_{n + 1}$ *(Industry i's share of manufacturing in country *j* in the initial period)

 $+ \alpha$ (Aid to country j*Sensitivity of industry i to exchange

rate appreciation) + ε_{ii} (1)

where *Growth*_{ij} is the annual average rate of growth of value added of industry *i* in country *j* over a ten-year period, obtained by normalizing the growth in nominal value added by the GDP deflator; $\zeta_{1,...,m}$ are the coefficients of the country fixed effects; $\zeta_{m+1,...,n}$ are the coefficients of the industry fixed effects; ζ_{n+1} is the coefficient of the initial period share of industry *i* in total value added in country *j* (which controls for convergence-type effects); Aid to country *j* is the average aid-to-GDP ratio for that country over the sample period. The coefficient of interest for us is α . It captures an interaction between a country-specific aid variable and an industry's sensitivity to the exchange rate effects induced by foreign aid. We posit that countries that receive more aid should see a more negative impact in industrial sectors that are more sensitive, so that we would expect the coefficient α to be negative.

The chief advantage of this strategy is that by controlling for country and industry fixed effects, the problem of omitted variable bias or model specification, which seriously afflicts cross-country regressions, is diminished. Essentially, we are making predictions about within-country differences between industries based on an interaction between a country and industry characteristic. Moreover, because we focus on differences between manufacturing industries (rather than between, say, manufacturing and services industries), we can rule out factors that would keep manufacturing underdeveloped as explanations of our results—for these factors should not affect the differences between manufacturing industries.

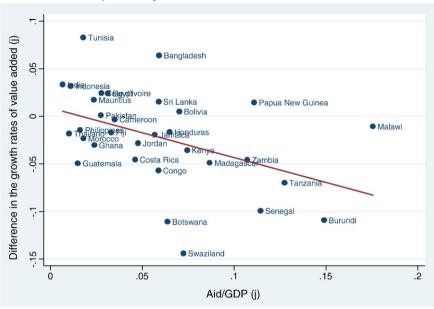
2.1. Implementing the strategy

However, this focus on manufacturing only (driven also by data availability) comes at a cost. The manufacturing sector, by and large, is

Table 2 Impact of aid on sectoral growth: core specification.

Dependent variable is annual average rate of growth of value added in industry <i>i</i> in country <i>j</i> .					
	OLS; 1980s (1)	OLS; 1980s (2)	OLS; 1990s (3)	OLS; 1990s (4)	
Initial industry share(ij)	-0.390^{***} [0.084]	-0.412^{***} [0.082]	-0.231 ^{**} [0.089]	-0.219 ^{***} [0.080]	
Aid/GDP(j)*exportability1 index(i)	-0.506^{**} [0.207]		-0.728 ^{**} [0.304]		
Aid/GDP(j)*exportability2		-1.042^{***}		-1.258^{***}	
index(i)		[0.311]		[0.485]	
Observations	684	684	357	357	
R-squared	0.415	0.424	0.328	0.335	

Estimates are based on the OLS procedure. All standard errors, reported below the coefficient estimates, are robust. ***, **, and * denote significance at 1, 5, and 10%, respectively. All equations include country and industry fixed effects. Initial industry share (*ij*) refers to the share of industry *i* in country *j* as a share of total manufacturing sector value added in country *j*. Aid/GDP (*j*) refers to the share of aid to GDP in country *j* averaged over the period. Exportability1 index is a dummy that takes on a value of 1 if an industry's ratio of exports to value added is greater than the median value, and 0 otherwise. Exportability2 index is a dummy that takes on a value of 1 for ISIC sectors 321–324, and 0 otherwise.



A: Based on Exportability1 Index

B: Based on Exportability2 Index

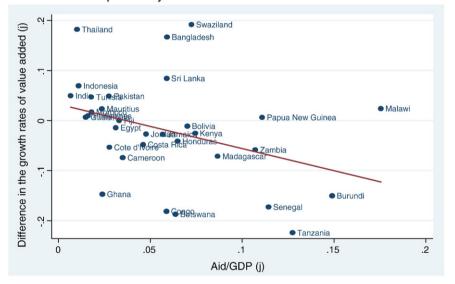


Fig. 3. Non-parametric depiction of core results. We divide the industries into two groups—exportable and others. In panel A, the exportable industries are as defined by the exportability1 index, and in panel B as defined by the exportability2 index. Next we estimate for each country the difference in average growth in value added between the exportable and other industries. The *y*-axis measures this difference, which is plotted against the aid-to-GDP received by each country (*x*-axis).

tradable. So how do we develop a proxy for the sensitivity, or relative sensitivity, of an industry to the competitiveness effects of aid?

One approach is to recognize that in a poor developing country, the lack of competitiveness is likely to show up to a greater extent in exports, than in import-competing industries. This is in part because, during the time period under study, in a poor country with an overvalued exchange rate, government actions to support import-competing industries (for example, through import tariffs and non-tariff barriers) were easier than actions to support exporting industries (for example, through cash or tax subsidies). Our proxy for exportability is an indicator variable for each industry, which takes the value 1 if the industry has a ratio of exports to value added (averaged across all countries in the sample) greater than the median across industries.⁵ The indicator takes the value zero otherwise. We

call this indicator "exportability1."⁶ We use another measure of exportability, which is simply a dummy equal to one for the four industries that have been most closely associated with the growth of developing countries as they have moved out of agriculture: textiles, clothing, leather, and footwear.Of course, not all exports are likely to be subject to adverse effects from aid. In particular, exports of extractable resources where labor intensity is low are unlikely to be affected by an overvalued real exchange rate. Fortunately, by focusing on manufacturing, we largely exclude such industries.

Because we examine growth differentials between industries within countries, the results are less sensitive to the rationale for why aid is given. For example, even if aid is given only to countries

⁵ We obtained these data from Nicita and Olarreaga (2001).

⁶ By discretizing our index we lose some information, but we also eliminate noise in classification. We also stay closer to the notion that either a product is exportable or it is not.

that display poor growth, inter-industry growth differentials should not be seriously affected. So, in our core specification, we will report our results using a simple OLS estimation strategy. However, there can be a potential, if indirect, problem of reverse causality. Suppose low growth is primarily because countries have overvalued exchange rates, and aid is systematically given to countries that have more overvalued exchange rates. In this case, we might be attributing to aid what is actually driven by trade and exchange rate policies. One way to address this is to correct directly for policies, which we attempt to do. Another is through instrumentation, which allows us to disentangle the direction of causality. We will, therefore, also show that our results are robust to instrumentation. Our instrumentation strategy, which draws directly from Rajan and Subramanian (2008), is described in the Appendix.

2.2. Data and their sources

The data and sources are described in detail in Appendix A. The data for industry value added growth come from the *Industrial Statistics Database* (2003) of the United Nations Industrial Development Organization (UNIDO). The data are at the 3-digit level of the International Standard Industrial Classification of All Economic Activities (ISIC, Revision 2) and are available for the 1980s and 1990s.

In order to keep the sample as large as possible without compromising our focus on long-term growth, we include all those industry observations where the average growth rate can be computed over at least a seven-year period in the decade. On this criterion, the UNIDO database has data for 47 developing countries for the 1980s and 31 countries for the 1990s.

But our methodology is most applicable when we include broadly similar countries with roughly similar levels of technological endowments. At the same time, we do not want to exclude recipients of significant aid. Our final sample therefore comprises countries that receive aid greater than 1% of GDP or are low-income countries according to World Bank definitions in the initial year of the sample. Our sample then has 32 countries for the 1980s and 15 countries for the 1990s. The UNIDO database contains data on 28 industries in these countries.⁷

In Table 1, we present means, medians, and standard deviations for the key variables in the analysis. The median growth rate of value added for industries is 1.7% for the decade 1980–90 and 3.7% for 1990–2000. The average aid inflow into the 32 countries in the 1980– 1990 sample is 5.3% of GDP and the average aid inflow into the 15 countries in the 1990–2000 sample is 4.7% of GDP. In much of the paper we will focus on the results for the 1980s because our sample is twice as large (32 countries and 684 observations) as it is for the 1990s (15 countries and 357 observations). We will only present some key results for the 1990s, and refer the reader to the working paper version for robustness checks.

Some additional concerns need to be discussed. First, given that a number of countries are not covered by the UNIDO database, the concern arises whether the countries in our sample differ from the typical aid recipient. Comparing the core 1980s sample that we use in this paper with a larger sample (comprising all countries selected on the same criteria as in this paper and for which manufacturing data are available in the World Bank's World Development Indicators), we find countries in our sample receive moderately less aid (5.6% relative to 7.0%) and grow somewhat faster (GDP growth 1.0% vs 0.7%, manufacturing growth 4.8% vs 3.6%). Countries in our sample are neither considerably worse in performance, nor indeed very different from the larger sample. The variation in our sample on these indicators is also not very different (see Appendix Table 1D). Prima facie, there is no clear evidence that our sample is "special."

A second potential concern, related to the above, is that only a few countries in our sample have significant manufacturing sectors, limiting the relevance of our analysis. On this, it is worth pointing out that the share of manufacturing in GDP in our sample is not very different from the larger set of developing countries that receive aid greater than 1% of GDP (see row 3 in Appendix Table 1D). In other words, it is possible that, and indeed this is what our core results purport to establish, this low share is a consequence of aid. Thus, low manufacturing share is not in itself a reason not to explore the impact of aid, although it is clearly relevant when evaluating claims about the overall growth impact of aid. We will, however, undertake some robustness tests to address this concern.

3. Aid and sensitivity to exchange rate appreciation

3.1. The core specification

We present our core specifications in Table 2 for the 1980s and 1990s corresponding to the relationship depicted as channel A in the schema in Section 2.1. The dependent variable is the annual average growth in value added in industry *i* in country *j*. We estimate the coefficient of the interaction between exportability and aid. The four columns correspond to our two measures of exportability for the 1980s and 1990s sample. In all four cases, the interaction term is negative and statistically significant at the 5% confidence level or above. Value added in exportable industries grows relatively more slowly than for other industries in countries that receive more aid.

What can we say about magnitudes? The estimate in column suggests that in countries that receive an extra 1 percentage point of aid, exportable industries (as defined by the exportability1 index) grow slower by about 0.5 percentage point per annum relative to non-exportable industries. The estimate in column 2 suggests that clothing and footwear sectors grow slower by about 1 percentage point a year relative to other industries in countries that receive 1 percentage point of GDP more in aid. This is quite substantial when compared with the average growth rate in the sample of 1.8%. The estimated magnitudes are slightly larger for the 1990s.

We depict this core result for the 1980s in a non-parametric form in Fig. 3. We divide the industries into two groups (above- and belowmedian) depending on their exportability. Next we estimate for each country the difference in average growth in annual value added between above-median and below-median exportable industries. We plot this difference against the aid-to-GDP received by each country. Fig. 3 shows that the difference in growth is negatively related to the aid received by a country, and no single country drives it.

3.2. Robustness to instrumentation

As we said earlier, there is no obvious reason to expect growth in particular industries to drive aid or to determine the inherent exportability of industries. However, there can be a potential, if indirect, problem of reverse causality. Suppose low growth is primarily because countries have overvalued exchange rates, and aid is systematically given to countries that have more overvalued exchange rates. In this case, we might be attributing to aid what is actually driven by trade and exchange rate policies. So, we check if our results are robust to instrumentation, using the instrument we developed in Rajan and Subramanian (2008).

The key idea for instrumentation in Rajan and Subramanian (2008) is to model the *supply* of aid based on donor-related rather than recipient-specific characteristics. In other words, we base our instrument on considerations that drive individual donors to give aid to a country other than those related to a country's level of income or growth. So, our construction of instruments starts from the bilateral donor–recipient relationship—including measures like colonial ties, common language, and relative size to predict bilateral aid flows—and

⁷ Appendix A lists all countries for which data were available as well as those countries that were included in the econometric analysis.

Impact of aid on sectoral growth: robustness to instrumentation.

	Manufacturing > 15%					
	(1)	(2)	(3)	(4)	(5)	(6)
Panel A: Dependent variable is annual average rate of g	growth of value added	in industry i in country	, j			
Aid/GDP(j)*Exportability1 index(i)	- 1.913***		-1.228^{**}		-1.561^{***}	
	[0.644]	0 = 1 0 1 1 1	[0.543]	1.001**	[0.543]	
Aid/GDP(j)*exportability2 index(i)		- 3.718***		- 1.381**		- 4.136***
Institutional quality * exportability1 index(i)		[1.180]		[0.693]	-0.053	[1.094]
institutional quality 'exportability'i index(i)					[0.074]	
Institutional quality * exportability 2 index(i)					[0.074]	-0.050
······································						[0.128]
Observations	684	684	350	350	622	622
R-squared	0.365	0.321	0.372	0.384	0.397	0.306
Panel B: First stage for IV in columns 1–6 of Panel A ab	ove					
Dependent variable is Aid/GDP(j)*exportability1 and A		ty2, respectively				
Fitted aid/GDP(j)*exportability1 index(i)	0.230***		0.298***		0.292***	
	[0.032]	0.271***	[0.037]	0 202***	[0.034]	0 200***
Fitted Aid/GDP(j)*exportability2 index(i)		[0.035]		0.302*** [0.039]		0.299*** [0.036]
Institutional quality*exportability1 index(i)		[0.055]		[0.055]	0.004	[0.050]
institutional quanty exportability index(r)					[0.012]	
Institutional quality*exportability2 index(i)						0.015
						[0.012]
Observations	684	684	350	350	622	622
R-squared	0.764	0.718	0.750	0.674	0.773	0.725

Estimates in this table use an instrumental variable (IV) procedure. Unless otherwise noted, all specifications are for the 1980s. All standard errors, reported below the coefficient estimates, are robust. ****, **, and * denote significance at 1, 5, and 10%, respectively. Standard errors are corrected, using the procedure in Frankel and Romer (1999), to allow for the fact that the instruments are estimated. All equations include country and industry fixed effects. Initial industry share (*ij*) refers to the share of industry *i* in country *j* as a share of total manufacturing sector value added in country *j*. Aid/CDP (*j*) refers to the share of aid to GDP in country *j* averaged over the period. Exportability1 index is a dummy that takes on a value of 1 if an industry's ratio of exports to value added is greater than the median value, and 0 otherwise. Exportability2 index is a dummy that takes on a value of 1 or ISIC sectors 321–324, and 0 otherwise. The specifications in columns 1, 2, 5 and 6 are based on the core sample, while those in columns 3 and 4 are based on a sample of countries whose manufacturing to GDP ratio exceeds 15%. In Panel B, the dependent variable (which is the endogenous regressor in Panel A) is the product of aid/GDP in country *j* times the relevant exportability index (*i*). Fitted aid is obtained from estimating a gravity-type model of bilateral aid flows as described in the Appendix. The institutional quality variable is from ICRGE (International Country Risk Guide) and is a measure of the quality of government institutions that affect property rights or the ability to conduct business. It is published by a private firm that provides consulting services to international investors.

this is then aggregated up to calculate predicted aid (as Frankel and Romer (1999) have done in the trade literature). This is in contrast to the literature that picks instruments directly at the level of the recipient country.

The results are presented in Table 3, columns 1 and 2. The secondstage results for the interaction terms are in Panel A, with the corresponding first stage in Panel B.⁸ The coefficient on the interaction term continues to be significant, and the magnitudes of these coefficients are greater than the OLS estimates. The results in Panel B show that there is no problem of weak instrumentation because the *F*-value for the instrument is above 50 in all columns, which comfortably exceeds the Staiger and Stock (1997) threshold of 10 for strong instruments.⁹

One concern might be that the countries in our base regression include ones with tiny manufacturing sectors, which may skew the results. As a check that this does not drive our results, we include only countries that had a manufacturing to GDP share greater than 15%. In Table 3, columns 3 and 4, we find the coefficient estimates for this narrower sample to also be statistically and economically significant.

Another concern is that the historical and colonial ties in our instrument might proxy for the quality of a country's institutions. From ICRGE (International Country Risk Guide), we obtain a measure of the quality of government institutions that affect property rights or the ability to conduct business. We then control for institutional quality (interacted with the exportability indicators to capture the fact that exportables may be more institution-intensive) in columns 5 and 6. The core interaction coefficient between aid and exportability remains qualitatively similar in magnitude and significance.¹⁰

4. The transmission mechanism from aid to sectoral growth

To summarize the results thus far, we have shown the link between aid and the relative growth of exportable industries. The theory we have laid out suggests this is because of the effect of aid on the real exchange rate, which in turn adversely affects the labor-intensive and exportable industries. This then leads to the natural question: how can we be sure that our core result does indeed reflect such an effect and what is the connection between aid and real appreciation? To answer these questions we can bring to bear more evidence, direct and indirect, relating to real exchange rate appreciation.

4.1. Does aid "cause" excess appreciation?

Before we elaborate on our results, we need to clarify the term "excess appreciation". Strictly speaking aid, via its spending effect, causes a currency appreciation which leads to a decline in exports. The problem, however, is that currency appreciation could also be a trend phenomenon associated with rising trade and incomes (the Balassa– Samuelson effect). Therefore, we want to pick up the measure of real

⁸ In this table, the standard errors in the second-stage regressions are corrected to take account of the fact that the instrument used in the first stage is estimated. We used the procedure in Frankel and Romer (1999) to do this correction.

⁹ With one endogenous variable, the *F*-value for the strength of the instrument is simply the square of the *t*-statistic associated with its coefficient.

¹⁰ In a longer version of this paper that is available online, we show that our results are robust to alternative samples, notably the exclusion of outliers, to clustering of standard errors, to alternative measures of value added growth, and to alternative measures and types of aid.

Impact of aid on sectoral growth: is real exchange rate the channel?

	Dependent variable is excess appreciation(j)	Dependent va	riable is annual a	verage rate of gro	wth of value add	ed in industry i i	n country j
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Aid/GDP(<i>j</i>) 1980s	3.175*** [0.280]						
Excess appreciation(j)*6	exportability1 index(i)	- 0.069*** [0.023]		- 0.058** [0.025]		-0.190*** [0.072]	
Excess appreciation(j)*e	exportability2 index(<i>i</i>)		-0.199*** [0.030]		-0.179*** [0.031]		-0.510*** [0.149]
Aid/GDP(<i>j</i>)* Exportabilit	ty1 index(<i>i</i>)			-0.289 [0.223]			
Aid/GDP(j)*exportability	y2 index (i)				-0.471 [0.313]		
Observations	674	674	674	674	674	674	674
R-squared	0.126	0.428	0.462	0.430	0.465	0.397	0.354
	in Columns 6–7 of Panel A above ess appreciation(j)*exportability1 and exportability2	respectively					
Initial industry share(<i>ij</i>)						0.034	0.032
Fitted aid/GDP(j)*expor	tability1 index(<i>i</i>)					[0.146] 1.972***	[0.106]
						[0.300]	
Fitted Aid/GDP(j)*expor	tability2 index(<i>i</i>)						2.158*** [0.318]
Observations						674	674
R-squared						0.538	0.239

Estimates are based on the OLS procedure, except for those in columns 6 and 7. All standard errors, reported below the coefficient estimates, are robust. ***, **, and * denote significance at 1, 5, and 10%, respectively. All estimations except for column (1) include country and industry fixed effects and initial industry share (*ij*) of industry *i* in country *j* as a share of total manufacturing sector value added in country *j*. Aid/GDP (*j*) refers to the share of aid to GDP in country *j* averaged over the period. Exportability1 index is a dummy that takes on a value of 1 if an industry's ratio of exports to value added is greater than the median value and 0 otherwise. Exportability2 index is a dummy that takes on a value of 1 for ISIC sectors 321–324, and 0 otherwise. The excess appreciation variable is from Johnson, Ostry and Subramanian (forthcoming). Note that the sample size is smaller than in the core specification. In columns 6 and 7, the instruments for excess appreciation are the same as those used for aid in Table 3.

appreciation that is over and above that suggested by the Balassa–Samuelson effect—that real exchange rates increase with real incomes. We follow Johnson, Ostry and Subramanian (forthcoming) (and implicitly, Dollar (1992), Easterly and Levine (2003), and Rodrik (2007)) in measuring excess appreciation or overvaluation.¹¹

In Table 4, we provide evidence that excess appreciation is the channel through which aid affects exports. In column 1, we show the simple correlation between excess appreciation and aid inflows which is statistically significant at the 1% level. If real appreciation is the channel, then like aid, it should particularly affect exportable sectors. In other words, in countries with more appreciated exchange rates, exportable and labor-intensive sectors should grow relatively slower than other sectors. We therefore estimate model (1) above with the difference that we replace the aid variable with our measure of excess appreciation. The results in columns 2 and 3 in Table 4 show that the interaction term between excess appreciation and either of our exportability measures is negative and significant. The coefficient estimates suggest that a one percentage point excess appreciation reduces the annual average growth of the exportable sectors by 0.07 and 0.2%, respectively, for the two measures of exportability. We have checked these results are robust to replacing our measure of excess appreciation with other measures of exchange rate overvaluation, including the recent work of Rodrik (2007) and Easterly and Levine (2003) and find that the results described here hold for those measures too (available from the authors upon request).

¹¹ The measure of excess appreciation is based on departures from long-run PPP. For every year in the sample period, we regress over the cross-section of countries, a country's price level of GDP from the Penn World Tables (6.1) on its real GDP per capita (in PPP terms). The deviation of the actual price level from the estimated price level is a measure of the country's excess appreciation. We averaged the annual values over the 1980s and 1990s respectively to obtain the final measure of excess appreciation. We have just established that excess exchange rate appreciation has a similar effect as aid on the growth in value added of exportable industries. Another way to check this is to introduce both the aid and excess appreciation interactions in the same regression, which we do in columns 4 and 5 of Table 4. If aid is indeed the deep causal determinant, and excess appreciation the main mediating channel, we should find that the direct effect of the aid interaction should be highly attenuated in the presence of the excess appreciation exportability interaction. Indeed, this is what we find. In both cases, the coefficient estimate of the excess appreciation interaction remains significant, while that for the aid interaction shrinks in both magnitude and statistical significance. This suggests that excess appreciation may represent the channel through which aid has influence.¹²

We offer one final piece of evidence in favor of excess appreciation as the channel, using our instrument for aid (interacted with exportability) in explaining the degree of excess appreciation (interacted with exportability) in the first stage, and use the predicted excess appreciation interaction in the second stage. The IV results for the second stage (columns 6 and 7, Panel A) suggest that the component of overvaluation caused by exogenous aid inflows does hurt competitiveness.

4.2. Is the excess appreciation a reflection of other policies rather than aid?

Is the observed significant estimate for the exportability-aid intensity interaction a proxy for other policies than aid? Consider trade reform. It is well-known that trade reform alleviates the antiexport bias of a regime. Is it possible that countries are poor because they have bad trade policies, and this, not aid, is responsible for the pattern of industry growth we observe? Indeed, could restrictive trade

¹² We obtained very similar results when we used the Easterly and Levine (2003) measure of overvaluation.

Impact of aid on sectoral growth: are there other channels?

Dependent variable is annual average rate of growth of value added in industry i in country j						
	(1)	(2)	(3)	(4)	(5)	
Panel A: Interacted with exportability1 index(i)						
Aid/GDP(j)*exportability1 index(i)	-0.547^{***}	-0.368^{*}	- 0.560**	-0.513**	-0.596***	
	[0.205]	[0.217]	[0.218]	[0.207]	[0.212]	
Policy(j) * exportability1 index(i)	-0.014					
	[0.024]					
Tariff*exportability1 index(i)		0.001**				
Dis also mentions and minute (i) * and a stability of index (i)		[0.000]	-0.003			
Black market premium(j)*exportability1 index(i)			[0.008]			
Standard deviation of inflation * exportability1 index(i)			[0.008]	0.000		
standard deviation of milation exportability mack(i)				[0.000]		
Aid/GDP (i) *financial dependence (i)				[]	0.316	
					[0.320]	
Observations	684	635	650	684	653	
R-squared	0.415	0.423	0.422	0.415	0.422	
Panel B: Interacted with exportability2 index(i)						
Aid/GDP(j)*Exportability2 index(i)	0.975***	0.818**	1.034***	1.043***	1.030***	
and a second	[0.324]	[0.331]	[0.318]	[0.312]	[0.320]	
Policy(j)*Exportability2 index(i)	0.023					
	[0.030]					
Tariff*Exportability2 index(i)		0.002*				
		[0.001]				
Black market Premium(j)*Exportability2 index(i)			0.004			
Standard deviation of inflation * Exportability2 index (i)			[0.007]	0.000		
Standard deviation of mination · Exportability2 mdex(1)				[0.000]		
Aid/GDP(i)*Financial dependence(i)				[0.000]	0.109	
					[0.323]	
Observations	684	635	631	684	653	
R-squared	0.424	0.438	0.431	0.424	0.429	

Estimates are based on the OLS procedure. All standard errors, reported below the coefficient estimates, are robust. ***, **, and * denote significance at 1, 5, and 10%, respectively. All equations include country and industry fixed effects and initial industry share (*ij*) of industry i in country j as a share of total manufacturing sector value added in country *j*. Aid/GDP (*j*) refers to the share of aid to GDP in country *j* averaged over the period. Exportability1 index is a dummy that takes on a value of 1 if an industry's ratio of exports to value added is greater than the median value, and 0 otherwise. Exportability2 index is a dummy that takes on a value of 1 for ISIC sectors 321-324, and 0 otherwise. The policy variable is from Wacziarg and Welch (2003), the tariff variable from the World Bank, the black market premium variable from Reinhart and Rogoff (2004), the inflation variability variable from the IMF, and the financial dependence variables from Rajan and Zingales (1998) (see Appendix Table 1A for details). Note also that the sample varies between columns because of data availability for some variables.

policies themselves be responsible for the observed excess appreciation (see, for example, Bhagwati and Dasai, 1970; Krueger, 1975) One way to test this is to include an interaction between the strength of trade liberalization policies and exportability. If the trade reform interaction swamps the aid interaction, we would have less confidence that aid causes the differential growth patterns we observe.

In Table 5A and B column 1, we include the Sachs–Warner measure of trade reform interacted with the exportability measures. We find that the coefficient on the trade policy interaction is insignificant but the aid interaction is relatively unchanged in magnitude and statistical significance. In column 2, we check whether trade policy as measured by the average level of tariffs interacted with exportability has any impact on the effect of the aid interaction. The aid interaction continues to have a negative and significant coefficient.¹³

One could argue that exchange rate mismanagement and distortions result in slow growth, aid inflows, and the observed relative growth patterns of industries. In column 3, we use a measure of the black market premium (from Reinhart and Rogoff (2004)) instead of the trade policy measure, and obtain similar results.¹⁴ In column 4, we introduce an interaction between macroeconomic instability (which we measure as the standard deviation of consumer price inflation) and exportability. Again, the aid-exportability coefficient estimate is relatively unchanged in these two specifications. Hence, the results in columns 1–4 suggests that it is aid rather than trade or macroeconomic policies that cause exchange rate appreciation, reinforcing the message obtained from instrumentation.

The next step is to show that aid does not affect industry growth through channels other than excess appreciation. One alternative explanation of the basic interaction between exportability and aid is that industries with a high need for capital (which need not be the exportable industries) grow relatively faster as a country receives aid inflows. This would be a relatively benign explanation of our basic findings, suggesting that aid relieves financing constraints and increases the overall resource envelope.¹⁵

There are three reasons why this is an unlikely explanation. First, the evidence in the introduction suggests the manufacturing sector (which is typically more capital-intensive than the services sector) is relatively smaller than the service sector in aid-receiving countries, not consistent with the benign "aid is financing" explanation. Second, we have also seen the adverse effects of aid on the relative growth rates of exportable industries come through an excess appreciation of the real exchange rate. Again, this is inconsistent with the benign explanation.

Third, if the capital-enhancing channel is at work (rather than the real-exchange-rate-excess appreciation channel), countries that

¹³ Tariffs are unweighted averages across all products and are obtained from the World Bank's database available (and explained in greater detail) at http://econ.worldbank.org/ WBSITE/EXTERNAL/EXTDEC/EXTRESEARCH/0,contentMDK:21051044~pagePK:64214825~ piPK:64214943~theSitePK:469382,00.html.

¹⁴ Rodriguez and Rodrik (1997) argue that both the Sachs-Warner and black market premium measures are measures not just of trade policy but of broader macroeconomic stability.

¹⁵ The simplest example of aid providing more resources to the private sector would be one where the government reduces its borrowing from the banking system in response to the aid, and hence makes more credit available to the private sector.

receive more aid should see an increase in the output of industries that are more dependent on external financial resources. To control for any possible effect of aid in alleviating financing constraints, we include the interaction of aid inflows with the Rajan and Zingales (1998) variable that measures the dependence of a particular industry on external resources to finance investment. Thus, if aid increases the resource envelope available to the industrial sector, we should expect the coefficient of this interaction term to be positive. Moreover, if the availability of capital rather than exportability is what matters, the coefficient on the aid-exportability interaction term should fall in magnitude when we include the aid-financial dependence interaction.

In Table 5A and B, column 5, we add another term, which is the interaction of aid and financial dependence of an industry. Only the aid-exportability interaction is statistically significant (and negative). The aid-financial dependence interaction is not statistically significant. Therefore, it does not appear that the channel through which the relative growth rates are affected by aid is via capital-intensive sectors getting increased access to resources. The magnitudes of the aid-exportability interaction coefficient are broadly similar to that estimated in columns 1 and 2 of Table 2, suggesting that exportability is not an indirect proxy for resource intensity.

5. Concluding remarks

Taken together, our results suggest there may indeed be an adverse impact of aid on the relative growth of exportable sectors, and that the channel through which these effects are felt is the exchange rate overvaluation induced by aid.

Despite the fact that for many aid-receiving countries, the manufacturing sector might be less important currently than agriculture, it is worth remembering that that was also true for many of the fast-growing countries when they first embarked upon development. Manufacturing exports provided the vehicle for their growth take-off, so any adverse effects on such exports should prima facie be a cause for concern about the effects of aid on growth.

The message from our work is that countries should avoid creating the conditions that generate uncompetitive exchange rates. To the extent that aid inflows are responsible, it would suggest a focus on the part of both donors and recipients on tailoring aid flows to the absorptive capacity for aid in the economy.

More generally, the theory does not suggest that the adverse effects of aid on competitiveness are inevitable, but our evidence indicates that the historical experience has not been positive. From a research perspective, it is perhaps more fruitful now to move beyond the inconclusive debate of whether aid is effective, and focus on specific ways it can be made to work better, by better understanding the reasons that might impair or enhance its effectiveness.

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Appendix A. Main data sources and description

- Industrial Statistics Database (2003) of the United Nations Industrial Development Organization (UNIDO) for data on value added and labor share. Data are at 3-digit level of the International Standard Industrial Classification of All Economic Activities (ISIC, Revision 2). [UNIDO database].
- World Development Indicators (World Bank) for the data on the share of manufacturing and services in GDP.
- WITS (World Integrated Trade Solution) data (World Bank) for exportability index.
- OECD's Development Assistance Committee (DAC) for data on aid.

A.1. Growth rate of real value added

The UNIDO dataset provides nominal value added both in terms of US dollars and local currency. The value added figure in US dollars is used for all regression analysis. The nominal value added (in current US dollars) was changed to a real value added (in constant Year 2000 US dollars), using the U.S. Producer Production Index provided by the International Monetary Fund's *International Financial Statistics* (IFS). This measure was, in turn, compared with the real value added in local currency to ascertain its reliability.¹⁶ More specifically, we required the correlation between the two be higher than 0.75 to be included in our base sample.

We then calculated the average annual growth rate of real value added for industry i in country j, for the 1980s and 1990s. We calculated this wherever data existed for at least a seven-year period.

Countries with aid less than 1% of GDP are not included in the regression analysis because they included a number of emerging market and other countries such as Malta, Cyprus, and Kuwait that cannot be considered meaningful aid recipients. Peru is dropped from the base sample due to its unusually high level of growth rates in all industries in the UNIDO database (i.e. exceeds 100% in all sectors). Niger is dropped from the 1990 sample as the data contained an observation where the ratio of wages to value added exceeded 17. We also dropped observations when this ratio exceeded one (this resulted in 10 and 12 observations being dropped respectively from the sample for the 1980s and 1990s).

Appendix B. Instrumentation strategy

Our instruments for aid are based on Rajan and Subramanian (2008). The key idea there is to model the *supply* of aid based on donor and donor–recipient relationships rather than on recipient-specific characteristics. In other words, we base our instrument on considerations that drive individual donors to give aid to a country other than those related to a country's level of income or growth. So, our construction of instruments starts from the bilateral (donor–recipient) relationship and aggregates up (as Frankel and Romer (1999) have done in the trade literature).

We posit that donors lend for reasons of history and in order to have influence. The greater the extent of historic relationships between a donor and a recipient the more likely that a donor will want to give aid. We capture historic relationships through colonial links and commonality of language.

One way to proxy for influence is through the relative size of donor and recipient. The bigger the donor is relative to the recipient, the

¹⁶ Since local PPI was not available for all developing countries in IFS, alternative deflators needed to be used to construct the measure of real value added in local currency. Accordingly, whenever PPI was not available, we used the effective deflator constructed with the index of industrial production as in Rajan and Zingales (1998). This deflator is the ratio of the growth rate of nominal value added in the entire manufacturing sector (from the UNIDO database) to the growth rate of the index of industrial production (from IFS). Alternatively, a GDP deflator was used whenever these two series were not available.

Appendix Table 1A

List of variables and data source.

Variable names	Description	Source
Growth rate of value added _{ij}	Industry i's annual growth rate of value added in country <i>j</i> , averaged over each decade.	UNIDO (2003)
Initial industry share _{ij}	Industry i's share in country j's total manufacturing value added at the beginning of the decade.	UNIDO (2003)
Exportability1 index _i	A dummy that takes a value of 1 if industry <i>i</i> has a ratio of exports to value that exceeds the industry median	WITS data, World Bank
	value. For each industry, the average ratio of exports to value added was calculated using a group of developing countries.	(at the 3-digit ISIC code). ^a
Exportability2 index _i	A dummy that takes a value of 1 for the four textiles and leather industries (ISIC 321–324).	WITS data, World Bank
		(at the 3-digit ISIC code).
Excess appreciation j	This measure is based on long-run PPP. For every year in the sample period, we regressed a country's price	Johnson, Ostry, and Subramanian
	level of GDP from the Penn World Tables (6.1) on its real GDP per capita (in PPP terms). The deviation of the	forthcoming
	actual price level from the estimated price level is a measure of the country's overvaluation. We averaged	
	the annual values over the 1980s and 1990s respectively to obtain the final measure of overvaluation	
Financial dependence _i	The measure of external financial dependence for all firms in industry <i>i</i> during the 1980s.	Rajan and Zingales (1998)
Aid/GDP _j	The ratio of aid-to-GDP for country j.	OECD DAC database.
Policy _j	Percent of years that a country is considered open.	Wacziarg and Welch (2003)
Tariffs	Unweighted average (across products) over years for which data are available	World Bank (website described in
		footnote 12 in the text.
Black market premium _j	Percent difference between a country's parallel market and official exchange rate expressed in terms of the latter.	Reinhart and Rogoff (2004)
Institutional quality _j	Several ICRGE indices averaged for the relevant period.	Bosworth and Collins (2003)
Life expectancy _j	Life expectancy at birth for beginning of the relevant decade.	WDI
Geography _j	Average of number of frost days and tropical land area.	Bosworth and Collins (2003)

^a The Trade and Production Database provides the WITS trade data at the 3-digit ISIC code. This database is available at: www.worldbank.org/research/trade.

Appendix Table 1B

Aid as a percentage of GDP. (*denotes inclusion in core sample).

1980s		1990s	
Malawi*	17.56%	Tanzania*	13.92%
Burundi*	14.90%	Ethiopia*	10.83%
Senegal*	12.41%	Senegal*	10.85%
Tanzania*	11.71%	Bolivia*	7.88%
Papua New Guinea*	11.09%	Kenya*	6.73%
Zambia*	10.72%	Jordan*	6.32%
Madagascar*	8.67%	Cameroon*	4.65%
Sri Lanka*	8.20%	Sri Lanka*	3.57%
Kenya*	8.13%	Egypt*	3.29%
Swaziland*	7.24%	Morocco*	1.55%
Honduras*	6.48%	Philippines*	1.37%
Botswana*	6.37%	Tunisia*	1.32%
Bolivia*	6.15%	Mauritius*	1.13%
Bangladesh*	5.91%	Indonesia*	1.02%
Congo*	5.87%	Costa Rica	0.94%
Jamaica*	5.68%	Panama	0.94%
Costa Rica*	4.61%	Algeria	0.64%
Israel	4.03%	India*	0.56%
Mauritius*	3.59%	Russia	0.46%
Fiji*	3.30%	Cyprus	0.34%
Jordan*	3.15%	Uruguay	0.29%
Egypt*	2.96%	Malaysia	0.28%
Cote d'Ivoire*	2.77%	South Africa	0.27%
Pakistan*	2.75%	Chile	0.25%
Cameroon*	2.59%	Colombia	0.24%
Ghana*	2.38%	Oman	0.18%
Tunisia*	2.25%	Venezuela	0.05%
Morocco*	2.14%	Korea	0.04%
Philippines*	1.86%	Singapore	0.03%
Guatemala*	1.48%	Kuwait	0.02%
Indonesia*	1.17%	China (Hong Kong)	0.02%
Thailand*	1.01%		
Malta	0.99%		
Panama	0.99%		
Cyprus	0.92%		
India*	0.76%		
Barbados	0.73%		
Uruguay	0.37%		
Algeria	0.33%		
Chile	0.32%		
Colombia	0.30%		
Bahamas	0.19%		
Singapore	0.16%		
Korea	0.08%		
Venezuela	0.07%		
Iran	0.06%		
Kuwait	0.03%		

more influence the donor is likely to have, so the ratio of the donor population to that of the recipient could be a good proxy for influence. In addition, a donor's influence is likely to be particularly pronounced if it is relatively large when it has close links with the recipient (so that it not only understands the pathways of influence but also potentially scares away other donors from seeking influence). We capture this channel by including the interaction between relative size and colonial links.

An example should help fix ideas; the United Kingdom should be willing to give more aid per capita to Uganda than to India; but it will be more willing to give aid to Uganda than to a similar-sized country

Appendix Table 1C Description of ISIC 3-digit industries.

ISIC code	Industrial sectors	Exportability1 index	Exportability2 index
311	Food products	1	0
313	Beverages	0	0
314	Tobacco	0	0
321	Textiles	1	1
322	Wearing apparel, except footwear	1	1
323	Leather products	1	1
324	Footwear, except rubber or plastic	1	1
331	Wood products, except furniture	1	0
332	Furniture, except metal	0	0
341	Paper and paper products	0	0
342	Printing and publishing	0	0
351	Industrial chemicals	1	0
352	Other chemicals	0	0
353	Petroleum refineries	1	0
354	Misc. petroleum and coal products	1	0
355	Rubber products	0	0
356	Plastic products	0	0
361	Pottery, china, earthenware	0	0
362	Glass and glass products	0	0
369	Other non-metallic mineral products	0	0
371	Iron and steel	0	0
372	Non-ferrous metals	1	0
381	Fabricated metal products	0	0
382	Machinery, except electrical	1	0
383	Machinery, electrical	0	0
384	Transport equipment	1	0
385	Professional and scientific equipment	1	0
390	Other manufactured products	1	0

Appendix Table 1D

Is the Core Sample Representative?

Variable	Sample in this paper	Larger sample
Aid to GDP		
Mean	5.3%	7.0%
Standard deviation	3.9%	5.6%
Real GDP growth		
Mean	1.0%	0.7%
Standard deviation	2.3%	2.1%
Real value added growth in manufacturing		
Mean	4.8%	3.6%
Standard deviation	3.5%	3.6%

Core sample comprises the 32 countries for which UNIDO's disaggregate manufacturing data are available and that have aid to GDP less than 1% or fall in the category of low-income countries. Larger sample includes countries that satisfy the same criteria but for which data on aggregate manufacturing is available in the World Bank's World Development Indicators (it excludes 3 countries with an aid-to-GDP ratio exceeding 30%).

in Africa, say Senegal, because France is likely to have a significant aid presence in the latter, thus diluting any possibility of British influence.

The aid supply decision from a donor (d) to a recipient (r) can be expressed by the following model:

$$\begin{split} A_{drt} / GDP_{rt} &= \beta 1' Y_{drt} + \upsilon_{drt} \\ &= \beta_0 + \beta_1 COMLANG_{dr} + \beta_2 CURCOL_{dr} + \beta_3 COMCOL_{dr} \\ &+ \beta_4 COMCOLUK_{dr} + \beta_5 COMCOLFRA_{dr} + \beta_6 COMCOLSPA_{dr} \\ &+ \beta_7 COMCOLPOR_{dr} + \beta_8 \log(POP_d / POP_r) \\ &+ \beta_9 \log(POP_d / POP_r)^* COMCOL_{dr} \tag{1} \\ &+ \beta_{10} \log(POP_d / POP_r)^* COMCOLFRA_{dr} \\ &+ \beta_{12} \log(POP_d / POP_r)^* COMCOLSPA_{dr} \\ &+ \beta_{13} \log(POP_d / POP_r)^* COMCOLPOR_{dr} + \upsilon_{drt} \end{split}$$

where A_{drt} is the aid given by donor *d* to recipient *r* in time period *t*. GDP_{rt} is the recipient country's GDP. The first three RHS variables capture historic factors: COMLANG is a dummy that takes a value of one if the donor and recipient share a common language; CURCOL is a dummy that takes a value of one if the recipient is *currently* a colony of the donor. COMCOL is a dummy that takes a value of one if the recipient was ever a colony of the donor. The next four variables simply disaggregate the colonial variable to capture difference across donors in the susceptibility of their giving to colonial ties (COMCOLUK, COMCOLFRA, COMCOLSPA, and COMCOLPOR refer in turn to colonial relationships involving respectively France, Portugal, Spain, and the United Kingdom).

The next set of variables relate to relative size. Donor size relative to the recipient is measured by the ratio of the logarithm of the populations of donor and recipient $(\log(POP_d/POP_r))$.¹⁷ Donor influence relative to other donors is additionally measured by the interaction of the relative population size and the colonial dummy $(\log(POP_d/POP_r) * COMCOL_{dr})$.

The estimated equation is then aggregated across donors to yield a level of the fitted value of aid-to-GDP for the recipient for that period.

$${\binom{A_{rt}}{GDP_{rt}}} = \sum_{d} \beta_1' Y_{drt}$$
(2)

In Rajan and Subramanian (2008), we present estimates for the model in (1) and show that they yield meaningful results. In Table 3 of Rajan and Subramanian (2008), we present estimates for the model

in (1). Virtually all the instrumenting variables are significant for all the time horizons, and between them the variables account for a reasonable share (between 33 and 42%) of the variation in the donor allocation decision. All the colonial relationships are significant as are the two measures of influence. The larger the donor relative to the recipient, the greater the aid given; and this effect is magnified, as conjectured, in cases where the donor had a colonial relationship with the recipient. For example, for the period 1960–00, a 1% increase in the ratio of donor size relative to recipient is associated with a 0.1% increase in the ratio of aid to GDP given by the average donor, an amount which is increased by 0.2% for a U.K. colony (which is the sum of 0.6% for the average colony minus 0.4% for the U.K. colony dummy), 1% for a Spanish colony; 1.9% for a French colony and 3.3% for a Portuguese colony. Thus, influence seems to matter a lot for donors, especially for Portugal and France.

A few caveats are worth noting about this instrumentation strategy which is not perfect. It is possible that Eq. (1) does not apply equally to all donor-recipient pairs. Our claim here is not to explain fully the aid allocation decision but to extract some (hopefully exogenous) information from that decision that would be reasonable for instrumentation. We would note that since our regressions are all cross-sectional, additional complications from the time series aspects of aid are not introduced. Finally, it should be noted that we do use a country-specific variable—namely population size—for instrumentation that might raise questions in terms of satisfying the exclusion restriction.

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¹⁷ To minimize endogeneity-related problems, we use the initial rather than the contemporaneous value of population.

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